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# Heterogeneous Networked Data Recovery from Compressive Measurements Using a Copula Prior

Nikos Deligiannis, *Member, IEEE*, João F. C. Mota, *Member, IEEE*, Evangelos Zimos,  
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**Abstract**—Large-scale data collection by means of wireless sensor network and internet-of-things technology poses various challenges in view of the limitations in transmission, computation, and energy resources of the associated wireless devices. Compressive data gathering based on compressed sensing has been proven a well-suited solution to the problem. Existing designs exploit the spatiotemporal correlations among data collected by a specific sensing modality. However, many applications, such as environmental monitoring, involve collecting heterogeneous data that are intrinsically correlated. In this study, we propose to leverage the correlation from multiple heterogeneous signals when recovering the data from compressive measurements. To this end, we propose a novel recovery algorithm—built upon belief-propagation principles—that leverages correlated information from multiple heterogeneous signals. To efficiently capture the statistical dependencies among diverse sensor data, the proposed algorithm uses the statistical model of copula functions. Experiments with heterogeneous air-pollution sensor measurements show that the proposed design provides significant performance improvements against state-of-the-art compressive data gathering and recovery schemes that use classical compressed sensing, compressed sensing with side information, and distributed compressed sensing.

**Index Terms**—Compressed sensing, side information, copula functions, air-pollution monitoring, wireless sensor networks.

## I. INTRODUCTION

THE emerging paradigm of smart cities has triggered the development of new application domains, such as environmental monitoring and smart mobility. These applications typically involve large-scale wireless sensor networks (WSNs) and internet-of-things (IoT) devices collecting and communicating massive amounts of environmental data, related to air pollution, temperature, and humidity. An air-pollution monitoring system<sup>1</sup>, for example, involves wireless devices spread

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<sup>1</sup>One can visit the websites of the European Environment Agency (EEA) (<http://www.eea.europa.eu/themes/air/air-quality>) and the USA Environmental Protection Agency (<https://www.epa.gov/aqs>).

in an urban area communicating measurements on several air pollutants, including carbon monoxide (CO), nitrogen dioxide (NO<sub>2</sub>), ozone (O<sub>3</sub>), and sulfur dioxide (SO<sub>2</sub>). Such data types have very different ranges and marginal statistics, but are intrinsically correlated.

This work shows how to effectively leverage the dependencies among diverse (alias, heterogeneous) data types in order to significantly reduce data-rates in the network. This reduction translates into power savings at the wireless nodes or IoT devices, which operate under austere limitations in energy resources. Efficient designs should, nevertheless, exploit intra- and inter-data dependencies at the decoder so as to conserve the computational effort at the wireless sensors and to diminish energy-demanding inter-sensor communication. Moreover, in order to safeguard power savings, devices should communicate over small distances through multi-hop wireless transmissions [3], namely, from neighbor to neighbor, rather than directly to a sink. Finally, as information is sent over error-prone wireless channels, data collection and recovery schemes should provide for robustness against communication noise.

## A. Prior Work

Related studies on the problem of data collection and recovery for WSNs proposed to reduce data rates by grouping nodes with correlated readings into clusters [4], [5] or by allowing a small subset of nodes to transmit data carrying most of the information in the network [6]. Alternative studies focused on conventional data compression techniques involving differential pulse-code modulation (DPCM) followed by entropy encoding [7], [8]. Other solutions considered collaborative wavelet transform coding [9] or offered a flexible selection between a distributed wavelet transform and a distributed prediction based scheme [10], [11]. These techniques, however, require additional inter-sensor communication, increasing the transmission of overhead information over the network.

An alternative strategy adheres to distributed source coding (DSC) [12], a paradigm that leverages inter-sensor (spatial) data correlation via joint decoding. DSC is a promising technique for WSNs as it shifts the computational burden towards the sink node and delivers code constructs that are robust against communication errors [12]. However, extending DSC to the multiterminal case is known to be a challenging problem in practice [13]–[15].

Compressed sensing (CS) [16], [17] addresses the problem of data aggregation in WSNs by enabling data to be recovered from a small set of linear measurements [18]. CS involves

solving an inverse problem at the decoder, for which several algorithms have been proposed, including orthogonal matching pursuit (OMP) [19], iterative thresholding [20], belief propagation (BP) [21], and approximate message passing (AMP) [22].

Considering a single-hop network, Haupt *et al.* [23] proposed CS-based data aggregation through synchronized amplitude-modulated transmissions of randomized sensor readings. Alternatively, Duarte *et al.* [24] proposed distributed compressed sensing (DCS), where random measurements are transmitted from each sensor and the data are jointly recovered at the sink by leveraging the spatiotemporal correlations. Furthermore, the authors of [25], [26] proposed a CS-based data aggregation method that used principal component analysis (PCA) to capture the spatiotemporal correlations in the data.

Assuming multi-hop transmission, Luo *et al.* [27] proposed a compressive data gathering method that alleviated the need for centralized control and complicated routing. They also presented measurement designs that limit the communication cost without jeopardising the data recovery performance. As an alternative solution, Lee *et al.* [28] proposed spatially-localized projection design by clustering neighboring nodes.

### B. Contributions

Prior studies on networked data aggregation via (distributed) compressed sensing [24]–[27] considered *homogeneous* data sources, namely, they proposed to leverage the spatiotemporal correlations within signals of the same type. Many applications, however, involve sensors of *heterogeneous* modalities measuring diverse yet correlated data (e.g., various air pollutants, temperature, or humidity). In this work, we propose a novel compressive data reconstruction method that exploits both intra- and inter-source dependencies, leading to significant performance improvements. Our specific contributions are as follows:

- We propose a new heterogeneous networked data recovery method, which builds upon the concept of Bayesian CS with belief propagation [21]. Our algorithm advances over this concept by incorporating multiple side-information signals, gleaned from heterogeneous correlated sources. This is in contrast to previous studies [29]–[32], which consider signal recovery aided by a single side information signal.
- Previous CS approaches describe the dependency among *homogeneous* sensor readings using the sparse common component plus innovations model [24]; simple additive models [33]; or joint Gaussian mixture models [32]. Unlike these studies, we model the dependency among *heterogeneous* data sources using copula functions [34], [35] and we explore copula-based graphical models—based on belief propagation [36]—for data recovery. Copula functions model the marginal distributions and the dependence structure among the data separately; as such, they capture complex dependencies among diverse data more accurately than existing approaches.
- Experimentation using synthetic data as well as diverse air-pollution sensor measurements from the USA Environmental Protection Agency [37] shows that, for a given

data rate, the proposed method reduces the reconstruction error of the recovered data with respect to classical CS [27], CS with side information [29], and DCS [24] based methods. Alternately, for a given reconstruction quality, the method offers significant rate savings, thereby resulting in less network traffic and reduced energy consumption at the wireless devices. Furthermore, the proposed design offers increased robustness against imperfections in the communication medium compared to the classical CS [27] and DCS [24] based methods.

### C. Outline

The paper continues as follows: Section II gives the background of the work and Section III details the proposed data recovery method. Section IV describes the copula-based statistical model for expressing the dependencies among diverse data types, whereas Section V elaborates on the proposed belief-propagation algorithm. Experimental results are provided in Section VI, whereas Section VII concludes the work.

## II. BACKGROUND

### A. Compressed Sensing

Compressed Sensing (CS) builds upon the fact that many signals  $\mathbf{x} \in \mathbb{R}^n$  have sparse representations, i.e., they can be written as  $\mathbf{x} = \Psi \mathbf{s}$ , where  $\Psi \in \mathbb{R}^{n \times n_0}$  is a dictionary matrix, and  $\mathbf{s} \in \mathbb{R}^{n_0}$  is a  $k$ -sparse vector (it has at most  $k$  nonzero entries). Suppose we observe  $m \ll n$  linear measurements from  $\mathbf{x}$ :  $\mathbf{y} = \Phi \mathbf{x} = \mathbf{A} \mathbf{s}$ , where  $\Phi \in \mathbb{R}^{m \times n}$  is a sensing (or encoding) matrix, and  $\mathbf{A} := \Phi \Psi$ . CS theory states that if  $\mathbf{A}$  satisfies the mutual coherence property [38], the Restricted Isometry Property [39], or the Null Space Property [40], then  $\mathbf{s}$  (and thus  $\mathbf{x}$ ) can be recovered by solving

$$\hat{\mathbf{s}} = \arg \min_{\mathbf{s}} \|\mathbf{s}\|_1 \quad \text{s.t.} \quad \mathbf{y} = \mathbf{A} \mathbf{s}. \quad (1)$$

In particular,  $\mathbf{s}$  is the only solution to (1) whenever the number of measurements  $m$  is sufficiently large. When the measurements are noisy, i.e.,  $\mathbf{y} = \mathbf{A} \mathbf{s} + \mathbf{z}$ , where  $\mathbf{z} \in \mathbb{R}^m$  represents additive noise,  $\mathbf{s}$  can be recovered by solving instead

$$\hat{\mathbf{s}} = \arg \min_{\mathbf{s}} \frac{1}{2} \|\mathbf{y} - \mathbf{A} \mathbf{s}\|_2^2 + \kappa \|\mathbf{s}\|_1, \quad (2)$$

where  $\kappa > 0$  controls the trade-off between sparsity and reconstruction fidelity. Instead of assuming that  $\mathbf{s}$  is strictly sparse (i.e.,  $\|\mathbf{s}\|_0 = k$ ), several works [21] (including this one) focus on compressible signals, i.e., signals whose coefficients decay exponentially, when sorted in order of decreasing magnitude.

### B. Compressed Sensing with Side Information

CS can be modified to leverage a signal correlated to the signal of interest, called side information, which is provided *a priori* to the decoder, in order to aid reconstruction [29]–[32], [41]. In CS with side information, the decoder aims to reconstruct  $\mathbf{x}$  from the measurements  $\mathbf{y}$ , the matrix  $\mathbf{A}$ , and a side information vector  $\mathbf{w}$  that is correlated with  $\mathbf{s}$ . The work in [29]–[31] provides guarantees for a particular way of

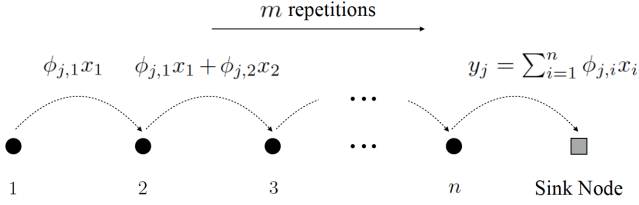


Fig. 1. Multi-hop transmission in a large-scale WSN using CS [27].

integrating side information into CS. In particular, one adds to the objective of (1) the  $\ell_1$ -norm of the difference between the optimization variable  $\mathbf{s}$  and the side information  $\mathbf{w}$ , yielding the  $\ell_1$ - $\ell_1$  minimization problem:

$$\hat{\mathbf{s}} = \arg \min_{\mathbf{s}} \|\mathbf{s}\|_1 + \|\mathbf{s} - \mathbf{w}\|_1 \quad \text{s.t.} \quad \mathbf{y} = \mathbf{A}\mathbf{s}. \quad (3)$$

Other studies considered prior information in the form of knowledge about the sparsity structure of  $\mathbf{s}$  [42]–[45] and derived sufficient conditions for exact reconstruction [42]. The authors of [46] proposed to recover the difference between the signal of interest and the side information, which was assumed to be sparser than the signal itself.

### C. Distributed Compressed Sensing

DCS [24] assumes a joint sparsity model to describe the spatiotemporal dependencies among  $\zeta$  homogeneous signals. The sensor signals  $\mathbf{x}_j \in \mathbb{R}^n, j \in \{1, 2, \dots, \zeta\}$ , are assumed to have a representation  $\mathbf{x}_j = \Psi(\mathbf{s}_c + \mathbf{s}_j)$ , where  $\mathbf{s}_c \in \mathbb{R}^n$  is a sparse component common to all signals,  $\mathbf{s}_j \in \mathbb{R}^n$  is a sparse innovation component unique to each signal, and  $\Psi \in \mathbb{R}^{n \times n}$  is the sparsifying basis. Each sensor  $j \in \{1, 2, \dots, \zeta\}$  independently encodes the measured signal by projecting it onto a sensing matrix  $\Phi_j$  and transmits the low-dimensional measurements  $\mathbf{y}_j = \Phi_j \mathbf{x}_j$  to the sink. The sink, in turn, jointly reconstructs the signals by solving:

$$\hat{\mathbf{s}}_{\text{all}} = \arg \min_{\mathbf{s}_{\text{all}}} \|\mathbf{s}_c\|_1 + \sum_{j=1}^{\zeta} \omega_j \|\mathbf{s}_j\|_1 \quad \text{s.t.} \quad \mathbf{y}_{\text{all}} = \mathbf{A}_{\text{all}} \mathbf{s}_{\text{all}},$$

where  $\omega_1, \dots, \omega_{\zeta} > 0$ ,  $\mathbf{y}_{\text{all}} = [\mathbf{y}_1^T \dots \mathbf{y}_{\zeta}^T]^T$  contains the measurements from all the sensors, and  $\mathbf{s}_{\text{all}} = [\mathbf{s}_c^T \mathbf{s}_1^T \dots \mathbf{s}_{\zeta}^T]^T$  the vector to be recovered, contains the common and all the innovation components. Also,

$$\mathbf{A}_{\text{all}} = \begin{bmatrix} \mathbf{A}_1 & \mathbf{A}_1 & \mathbf{0} & \mathbf{0} & \dots & \mathbf{0} \\ \mathbf{A}_2 & \mathbf{0} & \mathbf{A}_2 & \mathbf{0} & \dots & \mathbf{0} \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ \mathbf{A}_{\zeta} & \mathbf{0} & \mathbf{0} & \mathbf{0} & \dots & \mathbf{A}_{\zeta} \end{bmatrix},$$

where  $\mathbf{A}_j = \Phi_j \Psi$  is associated to sensor  $j \in \{1, 2, \dots, \zeta\}$ . Note that the  $j$ -th block equation of  $\mathbf{y}_{\text{all}} = \mathbf{A}_{\text{all}} \mathbf{s}_{\text{all}}$  corresponds to the measurements of sensor  $j$ :  $\mathbf{y}_j = \mathbf{A}_j(\mathbf{s}_c + \mathbf{s}_j)$ .

### D. Compressive Data Gathering for WSNs

The compressive data gathering approach in [18], [27] adheres to a multi-hop communication scenario in which

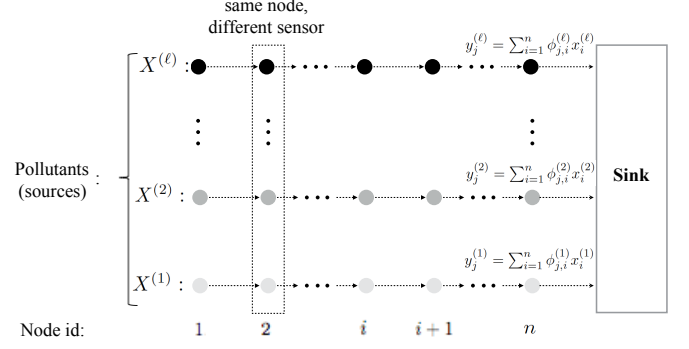


Fig. 2. Extension of the data gathering scheme of [27] to support the collection of diverse sensor data. For each source  $X^{(l)}$ ,  $l = 1, 2, \dots, \ell$ , the multi-hop transmission among the nodes takes place for  $m^{(l)}$  repetitions until the measurements vector  $\mathbf{y}^{(l)}$  is formed at the sink.

each node relays a weighted sum of sensor readings to a neighboring node. Specifically, consider a network of  $n$  nodes and let  $x_i \in \mathbb{R}$  denote a scalar reading of node  $i \in \{1, 2, \dots, n\}$ . As shown in Fig. 1, node 1 generates a pseudorandom number  $\phi_{j,1}$  — using its network address as the seed of a pseudorandom number generator — and transmits the value  $\phi_{j,1}x_1$  to node 2. Subsequently, node 2 generates  $\phi_{j,2}$ , computes the weighted sum  $\phi_{j,1}x_1 + \phi_{j,2}x_2$  and sends it to node 3. In sum, node  $k$  generates  $\phi_{j,k}$ , computes the value  $\phi_{j,k}x_k$ , adds it to the sum of the previous relayed values, and sends  $\sum_{i=1}^k \phi_{j,i}x_i$  to node  $k+1$ . The sink node thus receives  $y_j = \sum_{i=1}^n \phi_{j,i}x_i$ . After repeating the procedure  $m$  times, for  $j = 1, \dots, m$ , the sink obtains

$$\mathbf{y} = [\phi_1 \quad \dots \quad \phi_i \quad \dots \quad \phi_n] \mathbf{x} = \Phi \mathbf{x}, \quad (4)$$

where  $\mathbf{y} = (y_1, \dots, y_j, \dots, y_m)$  is the vector of measurements,  $\phi_i = (\phi_{1,i}, \dots, \phi_{j,i}, \dots, \phi_{m,i})$  is the column vector of pseudorandom numbers generated by node  $i$ , and  $\mathbf{x} = (x_1, \dots, x_i, \dots, x_n)$  is the vector of the node readings. Given the seed value and the addresses of the nodes, the sink can replicate  $\Phi$  and recover the data  $\mathbf{x}$  using standard CS recovery algorithms [19], [21], [22]. The study in [27] modified the sensing matrix in (4) as  $\Phi' = [\mathbf{I} \quad \mathbf{R}]$ , where  $\mathbf{I}$  is the  $m \times m$  identity matrix and  $\mathbf{R} \in \mathbb{R}^{m \times (n-m)}$  is a pseudorandom Gaussian matrix. This means that the first  $m$  nodes transmit their original readings directly to node  $m+1$ , which leads to a reduced number of transmissions in the network.

Alternatively, in the approach of [26], each node  $i$  transmits with probability  $p_i$  its reading directly to the sink. In this way, the sink collects measurements  $\mathbf{y} = \Phi \mathbf{x}$ , where  $\Phi$  is a very sparse binary matrix with one element equal to 1 per row and at most one element equal to 1 per column, while all the other elements are zero. The sink then solves (1) to recover the readings from all the nodes in the network.

## III. CS FOR HETEROGENEOUS NETWORKED DATA

State-of-the-art compressive data gathering and recovery solutions [18], [24], [26], [27] leverage the spatiotemporal correlation among *homogeneous* sensor readings, collected by a given sensing modality. However, current WSN and IoT setups

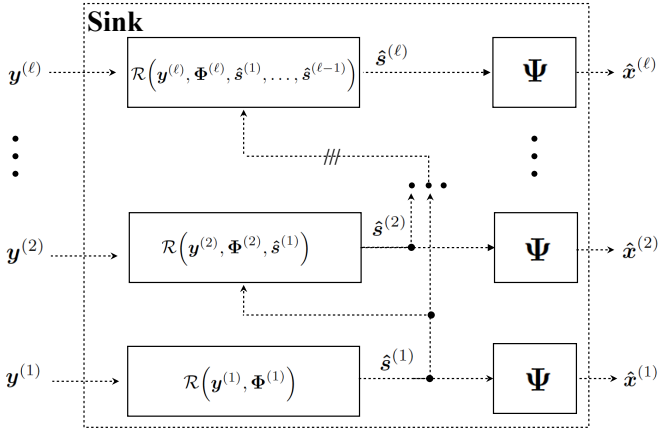


Fig. 3. Diagram of the proposed data recovery scheme. The vectors of readings  $\mathbf{x}^{(l)}$  of each data type are reconstructed sequentially,  $l = 1, \dots, \ell$ . The reconstruction of  $\mathbf{x}^{(l)}$  uses the respective measurements  $\mathbf{y}^{(l)}$  and matrix  $\Phi^{(l)}$ , as well as the sparse representations of the previously reconstructed modalities,  $\hat{\mathbf{s}}^{(1)}, \dots, \hat{\mathbf{s}}^{(l-1)}$ .

involve diverse sensing devices gathering *heterogeneous* data; for instance, different air pollution measurements (CO, NO<sub>2</sub>, O<sub>3</sub>, SO<sub>2</sub>) are collected in an environmental monitoring setup. We propose a design that jointly reconstructs heterogeneous correlated data from compressive measurements, by leveraging both intra- and inter-source data dependencies.

Consider a network comprising  $n$  wireless devices, each of which equipped with  $\ell$  sensors that monitor diverse, but statistically dependent, data types; for example, if the sensors measure the concentration of CO, NO<sub>2</sub>, O<sub>3</sub>, and SO<sub>2</sub>, then  $\ell = 4$ . Let  $x_i^{(l)}$  denote the reading at sensor  $i \in \{1, \dots, n\}$  of data type  $l \in \{1, \dots, \ell\}$ , and let  $\mathbf{x}^{(l)} = (x_1^{(l)}, \dots, x_i^{(l)}, \dots, x_n^{(l)})$  be the vector collecting all the readings of data type  $l$ . We assume that  $\mathbf{x}^{(l)}$  is sparse or compressible in a given orthonormal basis  $\Psi$ ; that is,  $\mathbf{x}^{(l)} = \Psi \mathbf{s}^{(l)}$ , where  $\mathbf{s}^{(l)} = (s_1^{(l)}, \dots, s_i^{(l)}, \dots, s_n^{(l)})$  is the compressible representation of  $\mathbf{x}^{(l)}$ . In our experiments in Section VI-B1,  $\Psi$  will be the discrete cosine transform (DCT) as, among several other common transforms, this is the one that yields the sparsest representation of pollution data.

The data gathering schemes [18], [26], [27] that were reviewed in Section II-D can be readily extended to address the collection of heterogeneous data. Fig. 2 shows how we modify the multi-hop scheme of [27]. Specifically, we assume that the communication network is a line graph, starting at node 1 and ending at node  $n$ . Node  $n$ , in turn, is connected to the sink node. The measurements of data type  $l$  are collected and transmitted as was described in Section II-D: node 1 measures  $x_1^{(l)}$ , and transmits  $\phi_{1,1}^{(l)} x_1^{(l)}$  to node 2, where  $\phi_{1,1}^{(l)}$  is randomly generated; node 2, in turn, measures  $x_2^{(l)}$ , generates  $\phi_{1,2}^{(l)}$ , computes  $\phi_{1,2}^{(l)} x_2^{(l)}$ , and transmits the sum  $\phi_{1,2}^{(l)} x_2^{(l)} + \phi_{1,1}^{(l)} x_1^{(l)}$  to node 3; and so on. The process is repeated  $m^{(l)}$  times, each time for different realizations of  $\phi_{j,i}^{(l)}$ . The sink then obtains the vector of measurements for source  $l$ :

$$\mathbf{y}^{(l)} = [\phi_1^{(l)} \dots \phi_i^{(l)} \dots \phi_n^{(l)}] \cdot \mathbf{x}^{(l)} = \Phi^{(l)} \mathbf{x}^{(l)}, \quad (5)$$

which has length  $m^{(l)}$ . Whenever the communication medium

and the receiver of the sink have imperfections, (5) can be modified to  $\mathbf{y}^{(l)} = \Phi^{(l)} \mathbf{x}^{(l)} + \mathbf{z}^{(l)}$ , where  $\mathbf{z}^{(l)} \in \mathbb{R}^{m^{(l)}}$  is additive white Gaussian noise (AWGN) [18]. The collection and transmission of measurements of the other data types is performed in the exact same way, either sequentially or concurrently.

Unlike the studies in [18], [26], [27], the sensing matrix we consider here is  $\Phi^{(l)} = \Theta^{(l)} \Psi^T$ , where  $\Theta^{(l)}$  is a sparse Rademacher matrix [21], and  $\Psi^T$  is the transpose of  $\Psi$ . Each measurement vector in (5) can then be written as  $\mathbf{y}^{(l)} = \Phi^{(l)} \mathbf{x}^{(l)} = \Theta^{(l)} \Psi^T \Psi \mathbf{s}^{(l)} = \Theta^{(l)} \mathbf{s}^{(l)}$ , where  $\Psi^T \Psi = \mathbf{I}$  because  $\Psi$  is orthonormal. Bearing a similarity with low-density parity-check matrices [47], sparse Rademacher matrices have only few non-zero entries, which are either  $-1$  or  $1$ , with equal probability. As shown in [21], they can lead to accurate and fast belief-propagation-based CS decoding, as opposed to dense Gaussian matrices [16], [17]. Similarly to the work in [21], the row weight  $\lambda$  and the column weight  $\rho$  of  $\Theta^{(l)}$  are kept very low—with respect to the dimension of the row and the column, respectively—and are assumed to be constant. Note that our selection for  $\Phi^{(l)}$  requires all nodes to know the matrices  $\Theta^{(l)}$  and  $\Psi^{(l)}$ , which can be accomplished by having all the nodes share a seed for generating the random entries of  $\Theta^{(l)}$ ; and, if required, the matrix  $\Psi^{(l)}$  can easily be pre-stored.

After receiving the measurements  $\mathbf{y}^{(l)}$  for all data types  $l = 1, \dots, \ell$ , the sink then proceeds to the data recovery stage, which is the focus of our paper, and is shown schematically in Fig. 3. Our method operates in stages, with the sink reconstructing the vectors  $\mathbf{s}^{(l)}$  sequentially, i.e., first  $\mathbf{s}^{(1)}$ , then  $\mathbf{s}^{(2)}$ , until  $\mathbf{s}^{(\ell)}$ . When reconstructing  $\mathbf{s}^{(l)}$ , the sink uses the measurements that were relayed,  $\mathbf{y}^{(l)}$ , the matrix  $\Phi^{(l)}$ , as well as the previously reconstructed vectors  $\hat{\mathbf{s}}^{(1)}, \dots, \hat{\mathbf{s}}^{(l-1)}$ , which play the role of multiple side information.

Standard CS recovery algorithms [19], [21], [22], as proposed by [18], [26], [27], would require recovering each sparse vector  $\mathbf{s}^{(l)}$  independently from the other vectors, based only on the measurement vector  $\mathbf{y}^{(l)}$ . This fails to leverage *inter-source* correlations. We will refer to this approach as the *baseline solution*. Alternatively, one can apply DCS [24] to recover the ensemble of sparse vectors  $\{\mathbf{s}^{(l)}\}_{l=1}^{\ell}$  using the ensemble of measurements vectors  $\{\mathbf{y}^{(l)}\}_{l=1}^{\ell}$  and the matrices  $\{\Phi^{(l)}\}_{l=1}^{\ell}$ . However, as shown in our experimental results, DCS does not efficiently capture the underlying dependencies among heterogeneous data, such as various air pollutants, which have different statistical properties.

The method we propose, in contrast, leverages diverse correlated signals through copula functions [34], [35]. Copula functions, explained in detail in Section IV-A, are elements of a statistical framework to effectively capture dependencies between random variables. As will be explained in Section V, we use copula functions to integrate knowledge from other data types in the reconstruction of a given data type or, in other words, as a way to integrate multiple side information. Our experiments in Section VI show that it is exactly because it uses multiple side information signals at the recovery stage that our scheme outperforms the state-of-the-art methods in [24], [27], [29]–[31].

#### IV. STATISTICAL MODELLING USING COPULAS

We now describe how to model statistically heterogeneous data using copula functions. Let  $S^{(l)}$  denote the random variable associated with the reading of source  $l \in \{1, \dots, \ell\}$ , and let  $s^{(l)}$  be one of its realizations. In general, the data sources we consider are not independent, meaning their joint probability density function (pdf)  $f_{S^{(1)}, S^{(2)}, \dots, S^{(\ell)}}(s^{(1)}, s^{(2)}, \dots, s^{(\ell)})$  does not factor into the product of its marginals. We will represent this joint pdf as  $f_{\mathbf{S}}(\mathbf{s})$ , where  $\mathbf{S} := (S^{(1)}, \dots, S^{(\ell)})$  is a random vector and  $\mathbf{s} := (s^{(1)}, \dots, s^{(\ell)})$  its realization.<sup>2</sup> We assume that each sensor  $i \in \{1, \dots, n\}$  observes  $\mathbf{S}_i$ , an independent realization of  $\mathbf{S}$ . In other words,  $\mathbf{S}_i$  is an i.i.d. copy of  $\mathbf{S}$ . This implies

$$\begin{aligned} f_{\mathbf{S}_1, \dots, \mathbf{S}_n}(\mathbf{s}_1, \dots, \mathbf{s}_n) &= \prod_{i=1}^n f_{\mathbf{S}_i}(\mathbf{s}_i) \\ &= \prod_{i=1}^n f_{S^{(1)}, \dots, S^{(\ell)}}(s_i^{(1)}, \dots, s_i^{(\ell)}). \end{aligned} \quad (6)$$

We will see next how a copula function enables working with the marginals of the joint pdfs  $f_{S^{(1)}, \dots, S^{(\ell)}}(s_i^{(1)}, \dots, s_i^{(\ell)})$  in (6), even though, as we saw before, these pdfs do not factor into the product of their marginals.

##### A. Introduction to Copulas

Suppose the random vector  $\mathbf{S} = (S^{(1)}, \dots, S^{(\ell)})$  is supported on a continuous set  $\mathcal{S} \subseteq \mathbb{R}^\ell$  and has joint cumulative distribution function (cdf)

$$F_{S^{(1)}, \dots, S^{(\ell)}}(s^{(1)}, \dots, s^{(\ell)}) = \Pr[S^{(1)} \leq s^{(1)}, \dots, S^{(\ell)} \leq s^{(\ell)}].$$

We will denote the marginal cdfs by  $F_{S^{(l)}}(s^{(l)}) = \Pr[S^{(l)} \leq s^{(l)}]$ . The probability integral transform [48] states that, independently of the distribution of  $S^{(l)}$ , the random variable  $U^{(l)} := F_{S^{(l)}}(S^{(l)})$  always has uniform distribution over  $[0, 1]$ .

The copula function of the random vector  $\mathbf{S} = (S^{(1)}, \dots, S^{(\ell)})$  is defined on the unit hypercube  $[0, 1]^\ell$  as the joint cdf of  $\mathbf{U} := (U^{(1)}, \dots, U^{(\ell)})$ , that is,

$$C(u^{(1)}, \dots, u^{(\ell)}) = \Pr[U^{(1)} \leq u^{(1)}, \dots, U^{(\ell)} \leq u^{(\ell)}], \quad (7)$$

where  $u^{(l)} = F_{S^{(l)}}(s^{(l)})$ . Namely, a copula is a multivariate cdf whose marginals have uniform distribution. The following result was seminal in the development of the theory of copula functions.

**Theorem 4.1 (Sklar's theorem [34]):** For any  $\ell$ -dimensional joint cdf  $F_{S^{(1)}, \dots, S^{(\ell)}}(s^{(1)}, \dots, s^{(\ell)})$  whose marginals are continuous, there exists a unique  $\ell$ -dimensional copula function  $C : [0, 1]^\ell \rightarrow [0, 1]$  such that

$$F_{S^{(1)}, \dots, S^{(\ell)}}(s^{(1)}, \dots, s^{(\ell)}) = C(u^{(1)}, \dots, u^{(\ell)}). \quad (8)$$

The implications of Theorem 4.1 are best seen after taking the  $\ell$ -th cross partial derivative of (8):

$$f_{S^{(1)}, \dots, S^{(\ell)}}(s^{(1)}, \dots, s^{(\ell)}) = \frac{\partial F_{S^{(1)}, \dots, S^{(\ell)}}(s^{(1)}, \dots, s^{(\ell)})}{\partial u^{(1)} \dots \partial u^{(\ell)}}$$

<sup>2</sup>Notice the difference in notation with respect to  $\mathbf{s}^{(l)} := (s_1^{(l)}, \dots, s_n^{(l)})$ , which collects the samples observed from data type  $l$ .

$$= c(u^{(1)}, \dots, u^{(\ell)}) \times \prod_{l=1}^{\ell} f_{S^{(l)}}(s^{(l)}), \quad (9)$$

where  $c(u^{(1)}, u^{(2)}, \dots, u^{(\ell)}) = \frac{\partial^\ell C(u^{(1)}, u^{(2)}, \dots, u^{(\ell)})}{\partial u^{(1)} \partial u^{(2)} \dots \partial u^{(\ell)}}$  denotes the copula density, and  $f_{S^{(l)}}(s^{(l)})$  is the pdf of  $S^{(l)}$ . Expression (9) tells us that the joint pdf of *dependent* random variables can be written as the product of the marginal pdfs, as if the variables were *independent*, times the copula density, which acts as a correction term. In other words, the copula density alone captures all the dependencies of the random variables. This means that finding a good model for the joint pdf boils down to finding not only accurate models for the marginal pdfs, but also an appropriate copula function to effectively capture the dependencies in the data.

##### B. Copula Families

There exist several bivariate and multivariate copula families [35], [49], [50], typically categorized into *implicit* and *explicit*. Implicit copulas have densities with no simple closed-form expression, but are derived from well known distributions. An example is the Elliptical copulas, which are associated to elliptical distributions (for example, the multivariate normal distribution), and have the advantage of providing symmetric densities. This makes them appropriate for high-dimensional distributions. Table I shows the expressions for the two mostly used Elliptical copulas: the Gaussian, and the Student's  $t$ -copula [34]. The expression for the Gaussian copula uses a standard multivariate normal distribution parameterized by the correlation matrix  $\mathbf{R}_G$ . In turn, the expression for the Student's  $t$ -copula uses a standard multivariate  $t$ -distribution, parameterized by the correlation matrix  $\mathbf{R}_t$  and by the degrees of freedom  $\nu$ . The diagonal entries of the correlation matrices  $\mathbf{R}_{(\cdot)}$  are 1, and the non-diagonal are equal to the estimated Spearman's  $\rho$  values.

Explicit copulas have densities with simple closed-form expressions but, being typically parameterized by few parameters, lack some modeling flexibility. The most popular explicit copulas are the Archimedean, which are parameterized by a single parameter  $\xi \in \Xi \subseteq \mathbb{R}$ . Specifically, an Archimedean copula is defined as [35]:

$$C_a(u^{(1)}, \dots, u^{(\ell)}; \xi) = q^{-1}\left(q(u^{(1)}; \xi) + \dots + q(u^{(\ell)}; \xi); \xi\right), \quad (10)$$

where  $q : [0, 1] \times \Xi \rightarrow [0, \infty)$  is a continuous, strictly decreasing, convex function such that  $q(1; \xi) = 0$ . The function  $q(u)$  is called generator and its pseudo-inverse, defined by

$$q^{-1}(u; \xi) = \begin{cases} q(u; \xi) & \text{if } 0 \leq u \leq q(0; \xi) \\ 0 & \text{if } q(0; \xi) \leq u \leq \infty, \end{cases} \quad (11)$$

has to be strictly-monotonic of order  $\ell$  [51]. Table II shows the distributions of the most popular Archimedean copulas: the Clayton, the Frank, and the Gumbel copulas [52].

For both families, the estimation of the copula parameters, e.g., the correlation matrix, is performed using training data. This will be described in detail in Section VI.

TABLE I  
ELLIPTICAL COPULA FUNCTIONS

Name	$C_e(u^{(1)}, \dots, u^{(\ell)})$	Parameters	Functions
Gaussian	$\Phi_{\mathbf{R}_G}(\Phi_g^{-1}(u^{(1)}), \dots, \Phi_g^{-1}(u^{(\ell)}))$	$\mathbf{R}_G$ : correlation matrix	$\Phi_{\mathbf{R}_G}$ : standard multivariate normal distribution $\Phi_g$ : standard univariate normal distribution
Student	$T_{\mathbf{R}_t, \nu}(t_\nu^{-1}(u^{(1)}), \dots, t_\nu^{-1}(u^{(\ell)}))$	$\mathbf{R}_t$ : correlation matrix $\nu$ : degrees of freedom	$T_{\mathbf{R}_t, \nu}$ : standard multivariate $t$ -distribution $T_\nu$ : univariate $t$ -distribution

TABLE II  
ARCHIMEDEAN COPULA FUNCTIONS

Name	$C_a(u^{(1)}, \dots, u^{(\ell)})$	Parameter Range $\Xi$	Generator $q(u)$
Clayton	$\left(\sum_{l=1}^{\ell} (u^{(l)})^{-\xi} - \ell + 1\right)^{-1/\xi}$	$\xi \in (0, \infty)$	$\xi^{-1} (u^{-\xi} - 1)$
Frank	$-\frac{1}{\xi} \log \left(1 + \frac{\prod_{l=1}^{\ell} (e^{-\xi u^{(l)}} - 1)}{(e^{-\xi} - 1)^{\ell-1}}\right)$	$\xi \in (-\infty, \infty)$	$-\log \left(\frac{e^{-\xi u} - 1}{e^{-\xi} - 1}\right)$
Gumbel	$\exp \left[ \left( -\sum_{l=1}^{\ell} (-\log u^{(l)})^{\xi} \right)^{1/\xi} \right]$	$\xi \in [1, \infty)$	$(-\log u)^{-\xi}$

### C. Marginal Statistics

As shown in (9), a consequence of Sklar's theorem (Theorem 4.1) is that copula functions enable us to work with the marginal pdfs of a random vector even when its components are not independent. We will consider the following pdfs when we model the distribution of each component.

#### 1) Laplace distribution

$$f_{S^{(l)}}(s^{(l)}; b^{(l)}) = \frac{1}{2b^{(l)}} \exp \left[ -\frac{|s^{(l)} - \mu^{(l)}|}{b^{(l)}} \right], \quad (12)$$

where  $b^{(l)}$  is the scaling parameter and  $\mu^{(l)}$  is the mean value for the  $l$ -th data type, with  $l \in \{1, 2, \dots, \ell\}$ .

#### 2) Cauchy (or Lorentz) distribution

$$f_{S^{(l)}}(s^{(l)}; \alpha^{(l)}, \beta^{(l)}) = \frac{1}{\pi \beta^{(l)}} \left[ 1 + \left( \frac{s^{(l)} - \alpha^{(l)}}{\beta^{(l)}} \right)^2 \right]^{-1}, \quad (13)$$

where  $\beta^{(l)}$  is a scale parameter specifying the half-width at half-maximum, and  $\alpha^{(l)}$  is the location parameter.

#### 3) Non-parametric distribution via kernel density estimation (KDE) [53]

$$f_{S^{(l)}}(s^{(l)}; h^{(l)}) = \frac{1}{n \cdot h^{(l)}} \sum_{i=1}^n \mathcal{K} \left( \frac{s^{(l)} - s_i^{(l)}}{h^{(l)}} \right), \quad (14)$$

where  $n$  is the number of samples from data type  $l \in \{1, 2, \dots, \ell\}$ . We use the Gaussian kernel  $\mathcal{K}(v) = \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}v^2)$  because of its simplicity and good fitting accuracy. We also select different smoothing parameters  $h^{(l)}$  for different data types,  $l \in \{1, \dots, \ell\}$ .

### V. COPULA-BASED BELIEF PROPAGATION

We now describe our reconstruction algorithm, executed at the sink node. As mentioned, the sparse vectors  $\mathbf{s}^{(l)}$  are reconstructed sequentially: first,  $\mathbf{s}^{(1)}$ , then  $\mathbf{s}^{(2)}$ , and so on. The reconstruction of each  $\mathbf{s}^{(l)}$  thus uses not only the respective measurements  $\mathbf{y}^{(l)}$ , but also the previously reconstructed data types  $\hat{\mathbf{s}}^{(1)}, \dots, \hat{\mathbf{s}}^{(l-1)}$  as side information.

We adopt the framework of Bayesian CS [21], [54], as it naturally handles our joint statistical characterization of the correlated modalities. We start by computing the posterior distribution of the random vector  $\mathbf{S}^{(l)}$ , representing the sparse vectors of coefficients of data type  $l$ , given the respective measurements  $\mathbf{Y}^{(l)}$  and the first  $l-1$  data types:

$$f_{\mathbf{S}^{(l)}|\mathbf{Y}^{(l)}\mathbf{S}^{(1)}\dots\mathbf{S}^{(l-1)}} \quad (15)$$

$$\propto f_{\mathbf{Y}^{(l)}|\mathbf{S}^{(1)}\dots\mathbf{S}^{(l)}} \times f_{\mathbf{S}^{(l)}|\mathbf{S}^{(1)}\dots\mathbf{S}^{(l-1)}} \quad (16)$$

$$= f_{\mathbf{Y}^{(l)}|\mathbf{S}^{(l)}} \times f_{\mathbf{S}^{(l)}|\mathbf{S}^{(1)}\dots\mathbf{S}^{(l-1)}} \quad (17)$$

$$= \prod_{j=1}^{m^{(l)}} f_{Y_j^{(l)}|\mathbf{S}^{(l)}} \times \prod_{i=1}^n f_{S_i^{(l)}|S_i^{(1)}\dots S_i^{(l-1)}}, \quad (18)$$

where we excluded the arguments of the pdfs for notational simplicity. From (15) to (16), we just applied Bayes's theorem and omitted constant terms. From (16) to (17), we used the assumption that measurements from data type  $l$  given realizations of all the previous data types  $j \leq l$  depend only on the value of  $\mathbf{S}^{(l)} = \mathbf{s}^{(l)}$ ; in other words, the process  $\mathbf{Y}^{(l)}|\mathbf{S}^{(1)}\dots\mathbf{S}^{(l)} = \mathbf{Y}^{(l)}|\mathbf{S}^{(l)}$  is Markovian. Finally, from (17) to (18), we used the assumption that measurement noise at different sensors is independent, and also that each sensor observes independent realizations of the random vector  $\mathbf{S} = (S^{(1)}, \dots, S^{(\ell)})$  (cf. Section IV). Obtaining an estimate of  $\mathbf{s}^{(l)}$  by minimizing the mean-squared-error or via maximum a posteriori (MAP) is challenging due to the complexity of the posterior distribution in (18). Therefore, as in [21], we use the belief propagation algorithm [36].

Our approach modifies the algorithm in [21] to take into account the previously reconstructed signals  $\mathbf{s}^{(1)}, \dots, \mathbf{s}^{(l-1)}$  in the reconstruction of  $\mathbf{s}^{(l)}$ . Fig. 4 represents the factor graph associated with (18). A factor graph represents the factorization of an expression by using two types of nodes: *variable nodes* and *factor nodes*. Variable nodes are associated to the variables of the expression, in this case, the components of the vector  $\mathbf{S}^{(l)} = (s_1^{(l)}, \dots, s_n^{(l)})$  and, in Fig. 4, are represented with circles. The factor nodes are associated to the intermediate



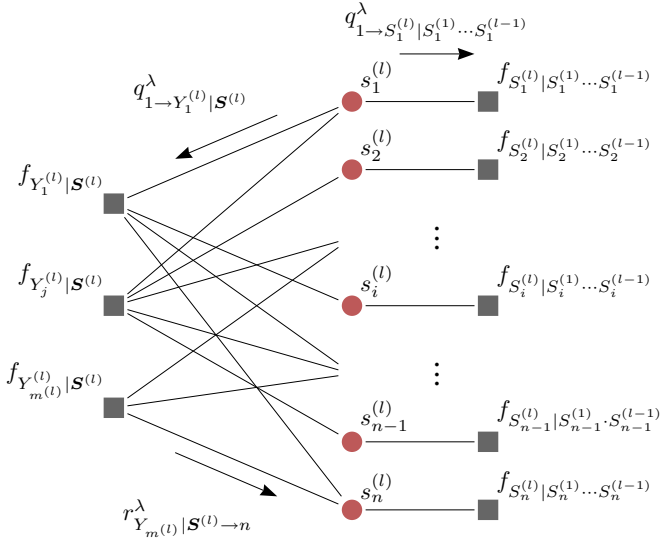


Fig. 4. Factor graph corresponding to the posterior distribution (18). The variable nodes are represented with circles and the factor nodes with squares. A message from variable node  $s_i^{(l)}$  to factor node  $f_Z$  at iteration  $\lambda$  is denoted with  $q_{i \rightarrow Z}^\lambda$ , and a message in the inverse direction is denoted with  $r_{Z \rightarrow i}^\lambda$ .

factors in the expression, in this case, the terms in (18) and, in Fig. 4, are represented with squares. Specifically, the leftmost squares in the figure represent the terms in the product  $\prod_{j=1}^{m^{(l)}} f_{Y_j^{(l)}|S^{(l)}}$ , and the rightmost squares represent terms in the product  $\prod_{i=1}^n f_{S_i^{(l)}|S_i^{(1)} \dots S_i^{(l-1)}}$ .

Notice that when there is no measurement noise in the acquisition of the measurements  $\mathbf{y}^{(l)}$ , each factor  $f_{Y_j^{(l)}|S^{(l)}}$  becomes

$$f_{Y_j^{(l)}|S^{(l)}}(\mathbf{y}_j^{(l)}|\mathbf{s}^{(l)}) = \delta\left(\mathbf{y}_j^{(l)} - \sum_{i=1}^N \Theta_{j,i}^{(l)} s_i^{(l)}\right), \quad (19)$$

where  $\delta(\cdot)$  is the Dirac delta function. When the measurement noise is, for example, AWGN, then  $\delta(\cdot)$  in (19) is replaced by the density of the normal distribution. Therefore, in Fig. 4, the edges from the variable nodes to the leftmost factor nodes represent the connections defined by measurement equation  $\mathbf{y}^{(l)} = \Theta^{(l)} \mathbf{s}^{(l)}$  (cf. Section III): there is an edge between factor  $f_{Y_j^{(l)}|S^{(l)}}$  and variable  $s_i^{(l)}$  whenever  $\Theta_{j,i}^{(l)} \neq 0$ . Recall also that the nonzero entries of  $\Theta^{(l)}$  are  $\pm 1$ .

Regarding the connections with the rightmost factor nodes in Fig. 4, notice that (9) implies that each term  $f_{S_i^{(l)}|S_i^{(1)} \dots S_i^{(l-1)}}$  can be expressed as the marginal pdf  $f_{S_i^{(l)}}$  times a correction term that captures information from the previously reconstructed data types. Indeed, assuming we have access to estimates  $\hat{\mathbf{s}}^{(k)}$  of  $\mathbf{s}^{(k)}$ , for  $k < l$ , there holds

$$f_{S_i^{(l)}|S_i^{(1)}, \dots, S_i^{(l-1)}}(s_i^{(l)} | \hat{s}_i^{(1)}, \dots, \hat{s}_i^{(l-1)}) \quad (20)$$

$$= \frac{f_{S_i^{(1)}, \dots, S_i^{(l)}}(s_i^{(1)}, \hat{s}_i^{(1)}, \dots, \hat{s}_i^{(l-1)})}{f_{S_i^{(1)}, \dots, S_i^{(l-1)}}(\hat{s}_i^{(1)}, \dots, \hat{s}_i^{(l-1)})} \quad (21)$$

$$= \frac{c(\hat{u}_i^{(1)}, \dots, \hat{u}_i^{(l-1)}, u_i^{(l)})}{c(\hat{u}_i^{(1)}, \dots, \hat{u}_i^{(l-1)})} \cdot f_{S_i^{(l)}}(s_i^{(l)}), \quad (22)$$

where  $\hat{u}_i^{(k)} = F_{S_i^{(k)}}(\hat{s}_i^{(k)})$  for  $k = 1, \dots, l-1$ . From (20) to (21) we used the definition of conditional density, and from (21) to (22) we simply used (9). Expression (22) depends only on  $s_i^{(l)}$  and thus explains the edges from the variable nodes to the rightmost factor nodes in Fig. 4.

Belief propagation is an iterative algorithm in which each variable node  $s_i^{(l)}$  sends a message to its neighbors  $\mathcal{M}_i$  (which are only factor nodes), and each factor node  $f_Z$  sends a message to its neighbors  $\mathcal{N}_Z$  (which are only variable nodes). Here,  $Z$  represents either  $Y_j^{(l)}|S^{(l)}$ , for  $j = 1, \dots, m^{(l)}$ , or  $S_i^{(l)}|S_i^{(1)} \dots S_i^{(l-1)}$ , for  $i = 1, \dots, n$ . In our case, a belief propagation message is a vector that discretizes a continuous probability distribution. For example, suppose the domain of the pdfs is  $\mathbb{R}$ , but we expect the values of the variables to be concentrated around 0. We can partition  $\mathbb{R}$  into 10 bins around 0, e.g.,  $(-\infty, -4] \cup (-4, -3] \cup \dots \cup (3, 4] \cup (4, +\infty)$ . The message, in this case, would be a 10-dimensional vector whose entries are the probabilities that a random variable belongs to the respective bin. For instance, all the messages to and from variable node  $s_1^{(l)}$  are vectors of probabilities,  $(\mathbb{P}\{S_1^{(l)} \in (-\infty, -4)\}, \dots, \mathbb{P}\{S_1^{(l)} \in (4, +\infty)\})$ , which are iteratively updated and represent our belief for the (discretized) pdf of  $s_1^{(l)}$ . Note, in particular, that all vectors have the same length and that all the messages to and from a variable node  $s_i^{(l)}$  depend on that variable only. We represent a message from variable  $s_i^{(l)}$  to factor  $f_Z$  at iteration  $\lambda$  as  $q_{i \rightarrow Z}^\lambda(s_i^{(l)})$ , and a message from factor  $f_Z$  to variable  $s_i^{(l)}$  as  $r_{Z \rightarrow i}^\lambda(s_i^{(l)})$ . The messages are updated as follows:<sup>3</sup>

$$q_{i \rightarrow Z}^\lambda(s_i^{(l)}) = \prod_{U \in \mathcal{M}_i \setminus \{Z\}} r_{U \rightarrow i}^{\lambda-1}(s_i^{(l)}) \quad (23)$$

$$r_{Z \rightarrow i}^\lambda(s_i^{(l)}) = \sum_{\sim s_i^{(l)}} f_Z(Z) \cdot \prod_{k \in \mathcal{N}_Z \setminus \{s_i^{(l)}\}} q_{k \rightarrow Z}^{\lambda-1}(s_k^{(l)}), \quad (24)$$

where  $\sum_{\sim s_i^{(l)}}$  denotes the sum over all variables but  $s_i^{(l)}$ , and a “product” between messages is the pointwise product between the respective vectors.

We run the message passing algorithm (23)-(24) for  $\Lambda$  iterations. To obtain the final estimate  $\hat{s}_i^{(l)}$  of each  $s_i^{(l)}$ , we first compute the vector

$$g(s_i^{(l)}) := \prod_{U \in \mathcal{M}_i} r_{U \rightarrow i}^{(\Lambda)}(s_i^{(l)}),$$

and select  $\hat{s}_i^{(l)}$  as the mid-value of the bin corresponding to the largest entry of  $g(s_i^{(l)})$ . This gives us each component of the estimated vector of coefficients  $\hat{\mathbf{s}}^{(l)}$ . In turn, the estimated readings are computed as  $\hat{\mathbf{x}}^{(l)} = \Psi \hat{\mathbf{s}}^{(l)}$ .

## VI. EXPERIMENTS

We evaluate the data recovery performance of the proposed copula-based design using synthetic data (cf. Section VI-A) as well as actual sensor readings taken from the air pollution database of the US Environmental Protection Agency

<sup>3</sup>See, e.g., [21] for a more detailed account on belief propagation algorithms, including a derivation of these formulas. Note also that, for simplicity, we omit normalizing constants.



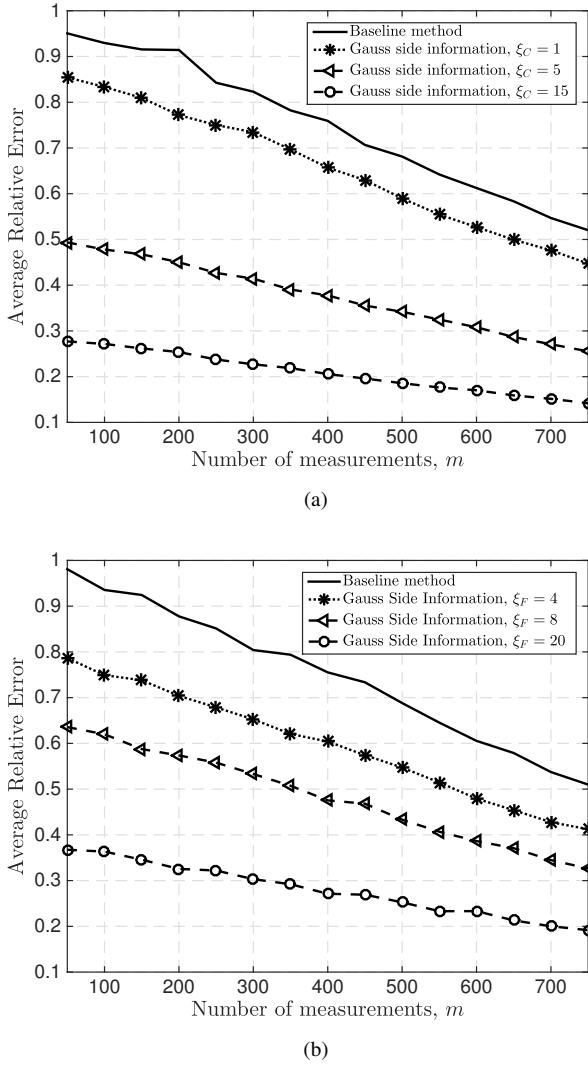


Fig. 5. Performance comparison of the proposed system against the baseline system using synthetic data. The marginal densities of the target and side information data follow the Laplace and Gaussian distribution, respectively. The generation of the data is done using (a) the Clayton or (b) the Frank copula function. The strength of the dependency is varied via controlling the  $\xi$  parameter of the copulas.

(EPA) [37] (cf. Section VI-B). Furthermore, in Section VI-C, we study the impact of the proposed method on the energy consumption of the wireless devices.

#### A. Results on Synthetic Data

In order to evaluate the proposed copula-based method, we simulate the approach described in Section III. We consider the vectorized readings  $\mathbf{x}^{(1)}, \mathbf{x}^{(2)} \in \mathbb{R}^{n \times 1}$  of two statistically dependent data types collected at a given time instance by a WSN and their compressible representations  $\mathbf{s}^{(1)}, \mathbf{s}^{(2)} \in \mathbb{R}^{n \times 1}$  in a basis  $\Psi$ . Following existing stochastic models [55] for the generation of spatially-correlated WSN data, we assume that both  $\mathbf{x}^{(1)}$  and  $\mathbf{x}^{(2)}$  are Gaussian. We also assume that  $\mathbf{x}^{(1)}$  is stationary (its variance is constant across readings), while  $\mathbf{x}^{(2)}$  is piece-wise stationary (its variance varies across groups of readings). Taking  $\Psi$  as the DCT basis, it can be fairly assumed that the coefficients in  $\mathbf{s}^{(1)}$  are Gaussian, whereas

the coefficients in  $\mathbf{s}^{(2)}$  follow the Laplace distribution<sup>4</sup> [56].

To simulate this scenario, we generate  $\mathbf{s}^{(1)}, \mathbf{s}^{(2)}$  as follows: We draw two coupled i.i.d. uniform random vectors  $\mathbf{u}^{(1)}, \mathbf{u}^{(2)}$ , with  $\mathbf{u}^{(l)} \in [0, 1]^n$ , from the bivariate Clayton or Frank copula [52]. The length of each uniform random vector is  $n = 1000$  and the copulas are parameterized by  $\xi_C$  and  $\xi_F$ , respectively. We consider different values for  $\xi_C = \{1, 5, 15\}$  and  $\xi_F = \{4, 8, 20\}$ , corresponding to weak, moderate, and strong dependency, respectively. We then generate the entries of  $\mathbf{s}^{(1)}$  by applying the inverse cdf of  $\mathcal{N}(0, \sigma^2)$  with  $\sigma = 4$  to the entries of  $\mathbf{u}^{(1)}$ ; similarly, the entries of  $\mathbf{s}^{(2)}$  are generated by applying the inverse cdf of  $\mathcal{L}(0, b)$  with  $b = 2$  to the entries of  $\mathbf{u}^{(2)}$ . We obtain measurements  $\mathbf{y}^{(2)} = \Theta^{(2)} \mathbf{s}^{(2)}$ —the column weight of  $\Theta^{(2)}$  is set to  $\rho = 20$ —and we assess the reconstruction of  $\mathbf{s}^{(2)}$ . We vary the number of measurements  $m^{(2)}$  from 50 to 750 and, for each  $m^{(2)}$ , we perform 50 independent trials—each with a different  $\Theta^{(2)}$  and  $\mathbf{s}^{(2)}$ —and we report the average relative error  $\|\mathbf{s}^{(2)} - \hat{\mathbf{s}}^{(2)}\|_2 / \|\mathbf{s}^{(2)}\|_2$  as a function of  $m^{(2)}$ .

We compare the recovery performance of two methods: the *baseline method*—which recovers  $\mathbf{s}^{(2)}$  from  $\mathbf{y}^{(2)}$  via Bayesian CS with belief propagation [21]—and the *proposed copula-based method* that recovers  $\mathbf{s}^{(2)}$  using  $\mathbf{y}^{(2)}$  and  $\mathbf{s}^{(1)}$ . In both methods, the length of each message vector carrying the pdf samples in the belief propagation algorithm is set to 243 and the number of iterations to 50. In order to have a fair comparison with CS, we account for a copula mismatch in our method. Namely, we use the bivariate Gaussian copula to model the dependency between the data, where the correlation matrix  $\mathbf{R}_G$  is fitted on the generated data using maximum likelihood estimation [57], even though the true relation between data types is generated with the Clayton or Frank copula.

The experimental results, depicted in Fig. 5, show that—despite the copula mismatch—the proposed algorithm manages to leverage the dependency among the diverse data and thus, to systematically improve the reconstruction performance compared to the classical method [21]. The performance improvements are increasing with the amount of dependency between the signals, reaching average relative error reductions of up to 72.90% and 64.09%, for the Clayton ( $\xi_C = 15$ ) and the Frank copula ( $\xi_F = 20$ ), respectively.

#### B. Results on Real Air Pollution Data

The AQS (Air Quality System) database of EPA [37] aggregates air quality measurements taken by more than 4000 monitoring stations, which collect hourly or daily measurements of the concentrations of six pollutants: ozone ( $\text{O}_3$ ), particulate matter (PM10 and PM2.5), carbon monoxide (CO), nitrogen dioxide ( $\text{NO}_2$ ), sulfur dioxide ( $\text{SO}_2$ ), and lead (Pb). We consider a network architecture comprising a sink and  $n = 1000$  nodes, where each node is equipped with  $\ell = 3$  sensors to measure the concentration of CO,  $\text{NO}_2$  and  $\text{SO}_2$  in the air. Using the node coordinates in the EPA database, we simulate such networks<sup>5</sup> by assuming that the transmission adheres to

<sup>4</sup>As shown in [56], the Laplace distribution emerges under the assumption that the variance across the group of readings is exponentially distributed.

<sup>5</sup>Each network is formed by nodes within only one of the following states: CA, NV, AZ, NC, SC, VA, WV, KY, TN, MA, RI, CT, NY, NJ, MD.

TABLE III  
AVERAGE PERCENTAGE OF THE NUMBER COEFFICIENTS OF THE DATA  
WITH AN ABSOLUTE VALUE BELOW A GIVEN THRESHOLD  $\tau$ .

	$\tau$	DCT	Haar	Daubechies-2	Daubechies-4
SO <sub>2</sub>	0.1	<b>28.8</b>	19.1	25.50	23.80
	0.2	<b>48.50</b>	35.80	46.30	44.00
	0.4	74.10	67.00	<b>77.50</b>	77.30
CO	0.1	<b>25.90</b>	16.40	22.60	19.40
	0.2	<b>44.50</b>	31.20	40.30	38.70
	0.4	<b>70.30</b>	58.20	69.80	69.20

LoRa [58], according to which the node distance does not exceed 2km in urban areas and 22km in rural areas. From the database, we take  $2 \times 10^5$  values for each of the three pollutants—i.e., CO, NO<sub>2</sub> and SO<sub>2</sub>—collected during the year 2015. The data are equally divided into a training and an evaluation set, without overlap.

1) *Sparsifying Basis Selection*: We first identified a good sparsifying basis  $\Psi$  for the data. Following the network architecture described in the previous paragraph, we organized the training data into blocks of  $n$  readings per pollutant. In order to form a block  $\mathbf{x}^{(l)}$ , readings must have the same timestamp and be measured by neighboring stations, adhering to the LoRa [58] transmission distance criteria. We projected the data in each block onto different set of bases, including the discrete cosine transform (DCT), the Haar, the Daubechies-2, and the Daubechies-4 continuous wavelet transform (CWT) bases; for the CWT we experimentally found that the scale parameter  $\alpha = 4$  led to the best compaction performance. Since the resulting representation  $\mathbf{s}^{(l)}$  is a compressible signal, we calculated the number of coefficients in  $\mathbf{s}^{(l)}$  whose the absolute value is below a given threshold  $\tau$ . Table III reports the results for SO<sub>2</sub> and CO, averaged over all the blocks in the training set. It shows that the DCT yielded the sparsest representations.

2) *Marginal Statistics and Copula Parameters*: To select the most appropriate marginal distribution for DCT coefficients of each  $\mathbf{s}^{(l)}$ , with  $l = 1, 2, 3$ , we performed fitting tests using the training set. The Laplace, the Cauchy, and the non-parametric distribution—via KDE with a Gaussian Kernel—were fitted to the data using the Kolmogorov-Smirnov test [59], with significance level set to 5%. The results, which were averaged over all the blocks in the training set, are reported in Table IV and Fig. 6. We can observe that the Cauchy distribution gives the best fit for the CO and SO<sub>2</sub> data, whereas the Laplace distribution best describes the statistics of the NO<sub>2</sub> data. The parameters of the distributions were estimated via maximum likelihood estimation (MLE), resulting in  $\hat{\beta}_{\text{CO}} = 0.6511$ ,  $\hat{\beta}_{\text{SO}_2} = 0.9476$  for the Cauchy distributions, and  $\hat{b}_{\text{NO}_2} = 2.3178$  for the Laplace distribution; recall the expressions for the pdf of these distributions in (12) and (13). We also estimated the mean values of the DCT coefficients, which were very close to zero for all distributions.

We now elaborate on the estimation of the parameters of the different copulas, described in Section IV-B. Using standard MLE [57], we calculate the correlation matrix  $\mathbf{R}_G$  for the Gaussian copula, the pairwise correlation values of which are presented in Table V. Moreover, we estimate the correlation

TABLE IV  
ASYMPTOTIC  $p$ -VALUES DURING THE KOLMOGOROV-SMIRNOV FITTING  
TESTS TO FIND THE MARGINAL DISTRIBUTION OF THE DCT  
COEFFICIENTS OF THE DATA.

	Laplace	Cauchy	KDE
CO	0.0031	0.6028	$9.4218 \times 10^{-20}$
NO <sub>2</sub>	0.5432	0.1441	$2.2777 \times 10^{-21}$
SO <sub>2</sub>	0.0471	0.9672	$1.0626 \times 10^{-21}$

TABLE V  
PAIRWISE COPULA PARAMETER ESTIMATES.

Parameters	(CO, NO <sub>2</sub> )	(NO <sub>2</sub> , SO <sub>2</sub> )	(CO, SO <sub>2</sub> )
Correlation	0.7025	0.8126	0.8563
Degrees of Freedom, $\nu$	35.56	35.56	490.95
$\xi$ (Clayton)	1.5770	2.3004	2.7655
$\xi$ (Frank)	6.6760	8.8767	11.0249
$\xi$ (Gumbel)	2.0877	2.5874	3.1619

matrix  $\mathbf{R}_t$  and the degrees-of-freedom parameter  $\nu$  for the Student's  $t$ -copula via approximate MLE [57]. The latter method fits a Student's  $t$ -copula by maximizing an objective function that approximates the profile log-likelihood for the degrees-of-freedom parameter. For the ensemble of the three pollutants we find the optimal value to be  $\nu = 89.91$ , whereas the values corresponding to each pair of pollutants are in Table V. Table V also reports the pair-wise maximum-likelihood estimates [60] of the  $\xi$  parameter for different bivariate Archimedean copulas [cf. (10)]. We consider bivariate Archimedean copulas for their simplicity, i.e., they are parameterized by a single parameter. This, however, limits their modeling capacity and makes them less accurate than, for example, Elliptical copulas [35].

3) *Performance Evaluation of the Proposed Algorithm*: We now describe how we evaluated the performance of our method against state-of-the-art reconstruction algorithms. Simulating the data collection approach described in Section III, for every vector of readings  $\mathbf{x}^{(l)}$  in the test dataset, we obtained its measurements as  $\mathbf{y}^{(l)} = \Phi^{(l)} \mathbf{x}^{(l)}$ . Similar to Section VI-A, we varied the number of measurements  $m^{(l)}$  from 50 to 750 and, for each  $m^{(l)}$ , we generated 50 different matrices  $\Phi$  (independently). We will report the average [over the  $\Phi$ 's and over all the points  $\mathbf{x}^{(l)}$  in the test dataset] relative error  $\|\mathbf{x}^{(l)} - \hat{\mathbf{x}}^{(l)}\|_2 / \|\mathbf{x}^{(l)}\|_2$  as a function of  $m^{(l)}$ .

In the first set of experiments, we used the NO<sub>2</sub> data to aid the reconstruction of the CO readings and considered the following methods: (i) the proposed copula-based belief propagation algorithm, running for 50 iterations and using five different bivariate copulas for modelling the joint distribution, namely, the Gaussian, the Student's  $t$ , the Clayton, the Frank, and the Gumbel copulas; (ii) the  $\ell_1$ - $\ell_1$  minimization method;<sup>6</sup> (iii) the baseline method [18], [27], which applies Bayesian CS [21] to recover the CO data independently; and, as a sanity check, (iv) simply keeping the  $m^{(l)}$  largest (in absolute value) DCT coefficients.

Fig. 7(a) depicts the relative reconstruction error versus the

<sup>6</sup>The  $\ell_1$ - $\ell_1$  minimization problem (3) is solved using the code in [61]; a detailed explanation of the solver can be found therein. The experiments in [29], [31] show that such a solver finds medium-accuracy solutions to (3) efficiently.

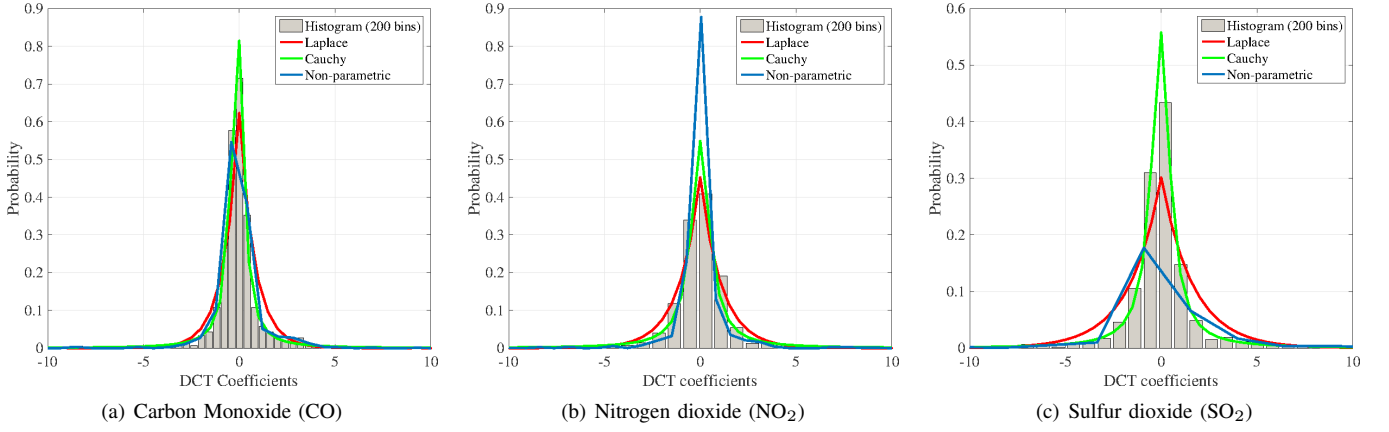


Fig. 6. Fitting of distributions (Laplace, Cauchy and KDE with a Gaussian kernel) on DCT coefficients of different air pollutants in the EPA dataset [37].

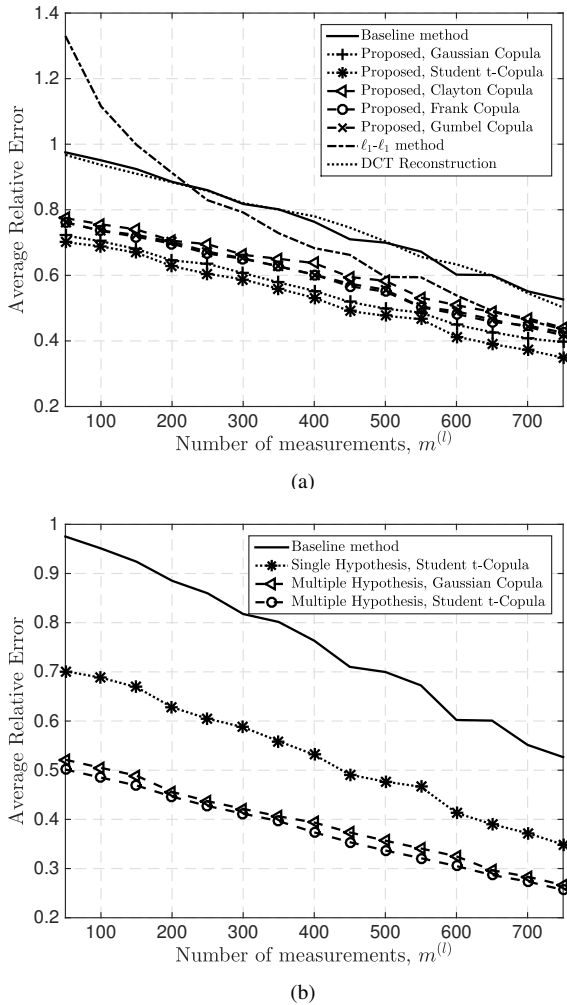


Fig. 7. Reconstruction performance of CO signals using as side information (a) only signals of  $\text{NO}_2$ , and (b) both signals of  $\text{NO}_2$  and  $\text{SO}_2$ . The baseline method refers to the no side information case, i.e., [18], [27].

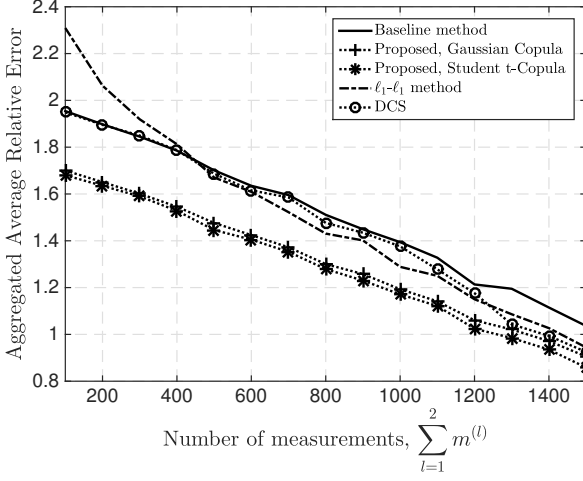
number of measurements  $m^{(l)}$ . It is clear that the proposed algorithm and  $\ell_1$ - $\ell_1$  minimization efficiently exploited the side information and were able to improve the performance with respect to the baseline method [18], [27]. When the number

of measurements was small ( $< 200$ ), the baseline method outperformed  $\ell_1$ - $\ell_1$  minimization; this is because, with few measurements, the side information was actually hindering reconstruction; recall that  $\ell_1$ - $\ell_1$  minimization assumes the side information to be of the same kind as the signal to reconstruct. Furthermore, it is clear that the proposed algorithm systematically outperformed  $\ell_1$ - $\ell_1$  minimization [29] for all the considered copula functions. The best performance was achieved by the Student  $t$ -copula function, providing average relative error reductions of up to 47.3% compared to  $\ell_1$ - $\ell_1$  minimization. We mention that, contrary to most results in compressed sensing, the results of Fig. 7(a) fail to exhibit a precise phase transition. This is because the representation of the data is not exactly sparse, only compressible. That can be seen in the plot, as the baseline method [18], [27] had a very similar performance to the DCT reconstruction, i.e., keeping only the largest DCT coefficients. This also shows that, in this case, what allowed both our method and  $\ell_1$ - $\ell_1$  minimization to achieve better performance was the proper use of the (correlated) side information.

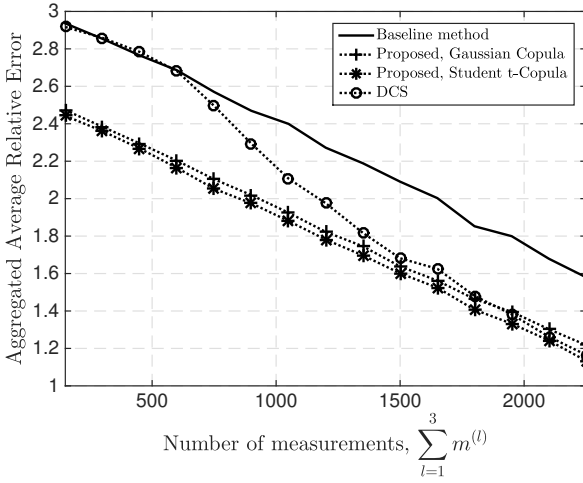
In another experiment, we reconstructed CO readings using as side information data from the other two pollutants, i.e.,  $\text{NO}_2$  and  $\text{SO}_2$ . Fig. 7(b) shows the average relative error of the proposed algorithm with one and two side information signals, and also the baseline method [18], [27]. It is clear that the more side information signals there are, the better the performance of our algorithm. We also observe that the Student's  $t$ -copula lead to a performance better than the Gaussian copula; this was because the former depends on more parameters than the latter, giving it a larger modeling capacity [62].

#### 4) Evaluation of the Aggregated System Performance:

We now describe the experiments conducted to evaluate the sequential reconstruction algorithm in which the readings are reconstructed consecutively. First, we focus on the scenario where two pollutants are measured, and we compare the following schemes: (i) the proposed sequential scheme, using the Gaussian and the Student  $t$ -copula models (as shown in Section VI-B3), they perform better than other copulas); (ii) sequential data recovery using  $\ell_1$ - $\ell_1$  minimization [29]; (iii)



(a)



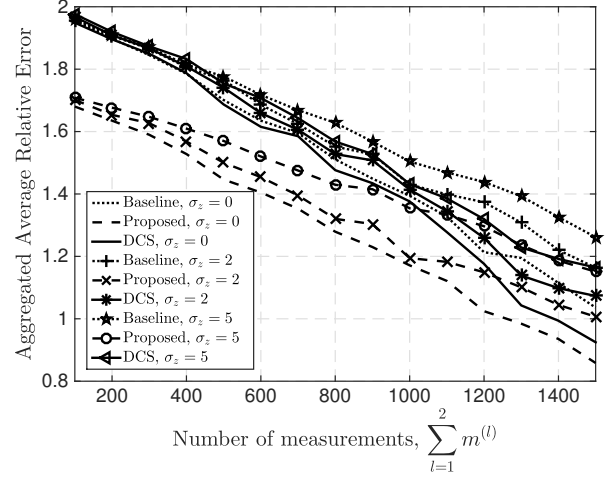
(b)

Fig. 8. Performance comparison of the proposed successive reconstruction architecture with the DCS, the ADMM-based and the baseline systems when we assume (a) two air pollutants (CO and NO<sub>2</sub>), and (b) three air pollutants (CO, NO<sub>2</sub> and SO<sub>2</sub>).

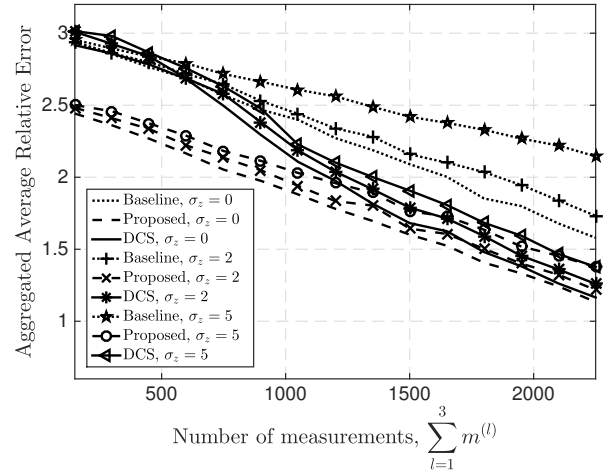
the DCS setup<sup>7</sup> [24]; and (iv) the baseline system in which each source is independently reconstructed using Bayesian CS [21].

The performance metric is expressed as the aggregated average relative error for all signals,  $\sum_{l=1}^L \|\mathbf{x}^{(l)} - \hat{\mathbf{x}}^{(l)}\|_2 / \|\mathbf{x}^{(l)}\|_2$ , versus the total number of measurements  $\sum_{l=1}^L m^{(l)}$ . Fig. 8(a) shows that the systems based on  $\ell_1$ - $\ell_1$  minimization and on DCS leverage both the inter- and intra-source dependencies between the pollutants, resulting in an improved performance with respect to the baseline system. However, when the number of measurements is small ( $\sum_{l=1}^2 m^{(l)} < 400$ ), we see that  $\ell_1$ - $\ell_1$  minimization performs poorly compared to the other methods. The proposed system with the Student  $t$ -copula model systematically outperforms all the other schemes, bring-

<sup>7</sup>In the classical DCS scenario, each signal of interest is constructed by many readings of the same sensor. In order to have a fair comparison with our design, we have modified this framework by assuming that each signal of interest contains readings from different sensors observing the same source. In our experiments we used  $\omega_1 = \dots = \omega_\ell = 1$ .



(a)



(b)

Fig. 9. Performance comparison of the proposed system with the DCS setup when we assume different noise levels ( $\sigma_z = 0, 2, 5$ ) for (a) two sources (CO and NO<sub>2</sub>), and (b) three sources (CO, NO<sub>2</sub> and SO<sub>2</sub>).

ing aggregated average relative error improvements of up to 27.2% and 13.8% against  $\ell_1$ - $\ell_1$  minimization and DCS, respectively.

When three pollutants are measured, we compared all the previous schemes, except the one based on  $\ell_1$ - $\ell_1$  minimization, since it does not handle multiple side information signals. Fig. 8(b) shows that DCS delivers superior performance compared to the baseline system, which is more noticeable when  $\sum_{l=1}^3 m^{(l)} > 600$ . Furthermore, the proposed design with the Student  $t$ -copula model provides significant aggregated average relative error reductions of up to 19.3% when compared to DCS [24]. It is important to notice that the proposed design significantly outperforms the other schemes when the number of measurements is small.

##### 5) Evaluation of the System Performance under Noise:

We evaluate the robustness of the proposed successive data recovery architecture against imperfections in the communication medium. As explained in Section III, we model such imperfections using a zero-mean white Gaussian noise

TABLE VI  
NUMBER OF MEASUREMENTS AND ENERGY CONSUMPTION AT THE WIRELESS NODES FOR TWO DIFFERENT DATA RECOVERY QUALITY LEVELS. TWO POLLUTANTS (CO, NO<sub>2</sub>) ARE MEASURED.

	Medium Data Recovery Quality			High Data Recovery Quality		
	Baseline	$\ell_1$ - $\ell_1$	Proposed	Baseline	$\ell_1$ - $\ell_1$	Proposed
Aggregated average relative error	1.4046	1.4230	1.3957	1.1148	1.1044	1.0969
Total number of measurements	950	850	550	1400	1300	1050
$E_{\text{proc.}}^{\text{HW}}$ (J)	$4.84 \times 10^{-6}$	$4.33 \times 10^{-6}$	$2.80 \times 10^{-6}$	$7.14 \times 10^{-6}$	$6.63 \times 10^{-6}$	$5.35 \times 10^{-6}$
$E_{\text{proc.}}^{\text{SW}}$ (J)	$46.51 \times 10^{-6}$	$41.62 \times 10^{-6}$	$26.93 \times 10^{-6}$	$68.54 \times 10^{-6}$	$63.64 \times 10^{-6}$	$51.41 \times 10^{-6}$
$E_{\text{Tx}}$ (J)	0.8586	0.7668	0.4968	1.2636	1.1718	0.9450
$E_{\text{total}}$ (J)	0.8586	0.7668	0.4968	1.2636	1.1718	0.9450

component  $\mathbf{z}^{(l)} \sim \mathcal{N}(\mathbf{0}, \sigma_z \mathbf{I})$  additive to the measurements, where  $\sigma_z$  is the noise standard deviation<sup>8</sup> and  $\mathbf{I}$  is the  $m^{(l)} \times m^{(l)}$  identity matrix. In this experiment, we vary the noise level as  $\sigma_z \in \{0, 2, 5\}$  and calculate the aggregated average relative error as a function of the total number of measurements. We first consider the case in which two pollutants are gathered by each device. The considered schemes are (i) the proposed system with successive data recovery using the copula-based algorithm (the Student's  $t$  copula is used); (ii) DCS [24], and; (iii) the baseline system [18], [27]. Figs. 9(a) and 9(b) show that the proposed system delivers superior performance compared to the competing systems for moderate ( $\sigma_z = 2$ ) and high ( $\sigma_z = 5$ ) noise. Moreover, we observe that the proposed algorithm is robust against noise, especially, when the number of measurements is small. In particular, the aggregated average relative error increases on average 5.8% ( $\sigma_z = 2$ ) and 18.7% ( $\sigma_z = 5$ ) with respect to the noiseless case.

In case three pollutants are measured, the proposed system systematically outperforms the DCS scheme and the baseline system, under both moderate and high noise. Moreover, the proposed design continues to demonstrate robustness against noise, with the aggregated average relative error increasing on average only 4.2% ( $\sigma_z = 2$ ) and 10.1% ( $\sigma_z = 5$ ) compared to the noiseless case. It is clear that the robustness of the proposed system increases with the number of pollutants.

### C. Energy Consumption Analysis

We now study the impact of the proposed system on reducing the number of measurements and in turn the energy consumption of the wireless nodes, for a given data reconstruction quality. The energy consumption at each node is broken down into a sensing, processing and transmission part:  $E_{\text{total}} = E_{\text{sens.}} + E_{\text{proc.}} + E_{\text{Tx}}$  [63]. The sensing part,  $E_{\text{sens.}}$ , depends on the amount of censored data; hence, its energy consumption is the same for the proposed and the baseline system. We thus focus our comparison on the energy consumption due to the processing and transmission parts. Following a typical IoT design, we assume that the nodes are equipped with the MSP430 micro-controller [64] and that communication adheres to LoRa [58]. MSP430 architectures [64] are typically built around a 16-bit CPU running at 25 MHz, with a voltage supply of  $V = 1.8 - 3.6$  Volt and a current of  $I = 200$

$\mu\text{A/MIPS}$  in the active mode. As discussed in Section III, every node generates a pseudorandom number, computes the product between this number and the censored value, adds it to the sum of the previous relayed values, and sends the final value to the next node. This operation is repeated per measurement and source  $l \in \{1, \dots, \ell\}$ . Neglecting the pseudorandom number generation part, the encoding operation boils down to a multiply-and-accumulate (MAC) operation. The MSP430 CPU cycles needed for a single signed 16-bit MAC operation are 17 for a hardware implementation or between 143 and 177 for a software implementation [64]. Therefore, the time to perform a single MAC operation<sup>9</sup> is  $t^{\text{HW}} = \frac{17 \text{ cycles}}{25 \text{ MHz}} = 0.68 \mu\text{s}$  or  $t^{\text{SW}} = \frac{177 \text{ cycles}}{25 \text{ MHz}} = 7.08 \mu\text{s}$ . The total time to encode the measurements at each device can then be calculated as  $t_{\text{proc.}}^{(\bullet)} = \sum_{l=1}^{\ell} m^{(l)} \times t^{(\bullet)}$  and the total processing energy as  $E_{\text{proc.}}^{(\bullet)} = V \cdot I \cdot t_{\text{proc.}}^{(\bullet)}$ , where  $(\bullet)$  stands for HW or SW. In order to calculate the energy consumption for transmission, we used the LoRa energy consumption calculator from Semtech [65], [66]. For a typical 12-byte payload packet with a 14 dBm power level, a current at 44 mA and a spreading factor of 7, the transmission energy consumption was estimated at 5.4 mJ. In the scenario where two pollutants (CO and NO<sub>2</sub>) are encoded [and no noise is assumed in the communication medium], Table VI-B4 reports the number of measurements and the energy consumption at the nodes for the baseline system, the system using  $\ell_1$ - $\ell_1$  minimization [29], and the proposed system. It is worth observing that the processing energy is negligible compared to the energy consumed by the transceiver. It is evident that for a comparable aggregated average relative error the proposed system leads to a significant reduction in the number of transmitted measurements compared to the competition, which translates to critical energy savings at the nodes.

## VII. CONCLUSION AND FUTURE WORK

We addressed the problem of data recovery from compressive measurements in large-scale WSN applications, such as air-pollution monitoring. In order to efficiently capture statistical dependencies among heterogeneous sensor data, we used copula functions [34], [35]. This enabled us to devise a novel CS-based reconstruction algorithm, built upon belief propagation [36], [67], which leverages multiple heterogeneous signals (e.g., air pollutants) as side information in order to improve reconstruction. Experiments using synthetic data

<sup>8</sup>We assume that the standard deviation of the noise is the same for all sources; hence, we drop the superscript ( $l$ ).

<sup>9</sup>We consider the higher value on the number of cycles for software.

and real sensor data from the USA EPA showed that the proposed scheme significantly improves the quality of data reconstruction with respect to prior state-of-the-art methods [23], [29], [68], even under sensing and communication noise. Furthermore, we showed that, for a given data reconstruction quality, the proposed scheme offers low encoding complexity and reduced radio transmissions compared to the state of the art, thereby leading to energy savings at the wireless devices. We conclude that our design effectively meets the demands of a large-scale monitoring application. Future work should concentrate on assessing the method on alternative datasets, such as the Intel-Berkeley Lab dataset [69], the dataset from the Center for Climatic Research [70], and the indoor dataset from the University of Padova [71].

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